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# Harris Geometric Distribution 

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#### Abstract

We introduce and characterize a new family of distributions, viz, Harris Geometric Distribution. The nature of hazard rate, entropy, distribution of minimum of sequence of i.i.d random variables is derived. Finally we test the goodness of fit with real data.


KEY WORDS: Geometric distribution, Marshall-Olkin family, Harris family, Failure rate.

## I. INTRODUCTION

The Harris geometric (HG) distribution is introduced by adding two extra parameters to geometric distribution with Harris distribution (Sandhya et al.[9],[10]). Marshall and Olkin [1] introduced a new method for adding a parameter to a family of distributions with application to the exponential and Weibull families. Marshall-Olkin Discrete Uniform (MODU) distribution has been developed by Sandhya and Prasanth [7] and this is found to be suitable for discrete data exhibiting either positive or negative skewness. Sandhya and Prasanth [8] have also considered Marshall-Olkin Geometric (MOG) distribution and some characterization of it. Prasanth and Sandhya[6] have developed Harris Discrete Uniform (HDU) distribution and studied its properties. Adding one or more parameters to a distribution makes it richer and more flexible for modeling data. In some discrete models, getting compact mathematical expressions for even simple descriptive statistics like expectation, variance and maximum likelihood estimates are difficult and here we make use of the numerical analysis based on the software package Mathematica to overcome such difficulties.

One of the main factors we consider here for characterization of the new family of distribution is the survival function ( s.f ). If $\mathrm{F}(\mathrm{x})$ is the s.f of a distribution, then by Marshall-Olkin method we get another s.f $\mathrm{G}(\mathrm{x})$, by adding a new parameter $\theta$ to it. That is, $\mathrm{G}(\mathrm{x}, \theta)=\theta \mathrm{F}(\mathrm{x}) /(1-(1-\theta) \mathrm{F}(\mathrm{x})),-\infty<\mathrm{x}<\infty, \theta>0$
(1) Let $g(x)$ is the probability mass function (p.m.f.) corresponding to $G(x)$, then the hazard rate of $X$ is,

$$
\begin{equation*}
\gamma_{\mathrm{G}}(\mathrm{x})=\mathrm{g}(\mathrm{x}) / \mathrm{G}(\mathrm{x}) \tag{2}
\end{equation*}
$$

A probability generating function (p.g.f.) considered by Harris [2] is defined as, $\mathrm{P}(\mathrm{s})=\mathrm{s} /\left(\mathrm{m}-(\mathrm{m}-1) \mathrm{s}^{\mathrm{k}}\right)^{1 / \mathrm{k}}, \mathrm{k} \in N$ where $N$ is the set of positive integers and $\mathrm{m}>1$. The p.g.f. of Harris distribution has been widely used in summation schemes. The role played by Harris distribution in schemes with random ( N ) sample sizes (random -sums or N -sums and random - extremes or N -extremes) in general and time series models in particular where N is a non-negative integer - valued r.v is discussed by satheesh et al. [12]. Satheesh and Sandhya [13] have been discussed this p.g.f. in the context of N -sums and N -extremes. Satheesh et a l.[12] discussed the stability of N-sums of r.v.s when N is Harris.

Adding a parameter by Marshall-Olkin technique for generalization of probability models has resulted very interesting changes in the performance of hazard rate function compared to that of base model. For example, the hazard rate of Marshall-Olkin extended exponential distribution can be constant, increasing as well as decreasing, where as the exponential distribution has only constant hazard rate. This has revealed the benefit that can be gained by adopting the extended model in reliability studies. Proceeding as the MO set up mentioned in (1) let $F(x)$ be any distribution function of a discrete r.v $X, f(x)$ be its p.m.f. and $F(x)$ be the reliability function of a distribution, then we can write the reliability function of a new distribution by substituting this $F(x)$ to get a new Harris family of s.f by adding a new parameter $\theta$ as, $\mathrm{H}(\mathrm{x}, \theta, \mathrm{k})=\left\{\theta \mathrm{F}^{\mathrm{k}}(\mathrm{x}) /\left[1-(1-\theta)\left(\mathrm{F}^{\mathrm{k}}(\mathrm{x})\right)\right]\right\}^{1 / \mathrm{k}} .0<\theta<\infty, \mathrm{k}>0$

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## II. HARRIS GEOMETRIC DISTRIBUTION

Consider a discrete random variable (r.v) X following geometric distribution with parameter ' p ' for $\mathrm{X}=1,2,3$ ...We know that the geometric distribution is the only distribution with lack of memory property. It is a discrete analog of the exponential distribution. It is the probability that the first occurrence of success require ' $x$ ' number of independent trials, each with success probability p . If the probability of success on each trial is $p$, then the probability that the $x^{\text {th }}$ trial (out of $x$ trials) is the first success is
$P(X=x)=(1-p)^{x-1} p \ldots \ldots$ with $(p+q=1)$, for $x=1,2,3 \ldots$ The probabilty distribution function is $F(x)=1-q^{x}$, the survival function (s.f) $\mathrm{F}(\mathrm{x})=\mathrm{q}^{\mathrm{x}}, \mathrm{E}(\mathrm{x})=1 / \mathrm{p}$ and Variance $\mathrm{q} / \mathrm{p}^{2}$, mode $=1$, Hazard rate $=\mathrm{p} / \mathrm{q}$, (which implies a constant failure rate).Then by adding a new parameter $\theta$, by (3) we get the s.f of the new distribution as,
$\mathrm{H}(\mathrm{x}, \theta, \mathrm{k})=\left\{\theta \mathrm{q}^{\mathrm{xk}} /\left[1-(1-\theta) \mathrm{q}^{\mathrm{xk}}\right]\right\}^{1 / \mathrm{k}}$.

$$
\begin{equation*}
=\theta^{1 / \mathrm{k}} \mathrm{q}^{\mathrm{x}} /\left[1-(1-\theta) \mathrm{q}^{\mathrm{xk}}\right]^{1 / \mathrm{k}}, 0 \leq \mathrm{q} \leq 1, \theta>0, \mathrm{k} \in N \text { where } N \text { is the set of positive integers } \tag{4}
\end{equation*}
$$

We call the new distribution with s.f (4) as Harris Geometric (HG) distribution and write, $X_{\sim} H G(q, \theta, k) .0 \leq q \leq 1, \theta>$ $0, \mathrm{k} \in N$ where $N$ is the set of positive integers

Note:
Throughout this paper we have, $0 \leq \mathrm{q} \leq 1, \theta>0$ and $k \in N$ where $N$ is the set of positive integers.
From (4) the p.m.f. of $X$ is, $h(q, \theta, k)=\left\{\theta^{1 / k} q^{x-1} /\left[1-(1-\theta) q^{(x-1) k}\right]^{1 / k}\right\}-\left\{\theta^{1 / k} q^{x} /\left[1-(1-\theta) q^{x k}\right]^{1 / k}\right\} \cdot x=1,2,3, \ldots$

$$
\begin{equation*}
=\theta^{1 / k} q^{x-1}\left\{\left[1 /\left(1-(1-\theta) q^{(x-1) k}\right)^{1 / k}\right]-\left[q /\left(1-(1-\theta) q^{\mathrm{xk}}\right)^{1 / k}\right]\right\} \cdot x=1,2,3, \ldots \tag{5}
\end{equation*}
$$

Remark 1 When $\mathrm{k}=1, \operatorname{HG}(\mathrm{q}, \theta, \mathrm{k})$ reduces to $\operatorname{MOG}(\mathrm{q}, \theta)$ with s.f $\mathrm{G}(\mathrm{q}, \theta)$ equal to $\theta \mathrm{q}^{\mathrm{x}} /\left[1-(1-\theta) \mathrm{q}^{\mathrm{x}}\right] . \mathrm{X}=1,2,3 \ldots \theta>0$.
Remark 2 If $\mathrm{H}(\mathrm{x}, \theta, \mathrm{k})$ is the s.f of $\mathrm{HG}(\mathrm{q}, \theta, \mathrm{k})$, then, $\mathrm{F}(\mathrm{x}) \leq \mathrm{H}(\mathrm{x}, \theta) \leq \theta^{1 / \mathrm{k}} \mathrm{F}(\mathrm{x})$ if $\theta>1$ and $\theta^{1 / k} \mathrm{~F}(\mathrm{x}) \leq \mathrm{H}(\mathrm{x}, \theta) \leq \mathrm{F}(\mathrm{x})$ if $0<\theta<1$.
Since the p.m.f. of the HG distribution is not in a compact form we numerically evaluate the characteristics associated with it. We compute the p.m.f. for different $q, k \& \theta$ (Table 1 and Table 2) plot the graph of the p.m.f.of X (Figure 1)

Table 1 The p.m.f. table of $\mathrm{X} \sim \mathrm{HG}(\mathrm{q}, \theta, \mathrm{k})$, for different k and $\theta$ with $\mathrm{q}=0.1$.

| X | $\theta=2 \mathrm{k}=5 \mathrm{q}=0.1$ | $\theta=0.5 \mathrm{k}=5$ <br> $\mathrm{q}=0.1$ | $\theta=0.2 \mathrm{k}=2$ <br> $\mathrm{q}=0.1$ | $\theta=2 \mathrm{k}=2$ <br> $\mathrm{q}=0.1$ | $\theta=0.5 \mathrm{k}=2$ <br> $\mathrm{q}=0.1$ | $\theta=5 \mathrm{k}=2$ <br> $\mathrm{q}=0.1$ | $\theta=10 \mathrm{k}=2$ <br> $\mathrm{q}=0.1$ |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 1 | 0.8851304 | 0.9129449 | 0.9550987 | 0.8592805 | 0.9291119 | 0.7807355 | 0.6971087 |
| 2 | 0.1033826 | 0.07834964 | 0.04042901 | 0.1265781 | 0.06381688 | 0.1969083 | 0.2712827 |
| 3 | 0.01033829 | 0.007834955 | 0.004025101 | 0.01272722 | 0.006364138 | 0.02012015 | 0.02844629 |
| 4 | 0.001033829 | 0.0007834955 | 0.000402492 | 0.001272791 | 0.000636396 | 0.002012457 | 0.00284603 |
| 5 | 0.00010338 | $7.834955 \mathrm{e}-05$ | $4.024922 \mathrm{e}-05$ | 0.000127279 | $6.363961 \mathrm{e}-05$ | 0.000201246 | 0.00028460 |
| 6 | $1.033829 \mathrm{e}-05$ | $7.834955 \mathrm{e}-06$ | $4.024922 \mathrm{e}-06$ | $1.272792 \mathrm{e}-05$ | $6.363961 \mathrm{e}-06$ | $2.012461 \mathrm{e}-05$ | $2.84605 \mathrm{e}-05$ |
| 7 | $1.033829 \mathrm{e}-06$ | $7.834955 \mathrm{e}-07$ | $4.024922 \mathrm{e}-07$ | $1.272792 \mathrm{e}-06$ | $6.363961 \mathrm{e}-07$ | $2.012461 \mathrm{e}-06$ | $2.84605 \mathrm{e}-06$ |
| 8 | $1.033829 \mathrm{e}-07$ | $7.834955 \mathrm{e}-08$ | $4.024922 \mathrm{e}-08$ | $1.272792 \mathrm{e}-07$ | $6.363961 \mathrm{e}-08$ | $2.012461 \mathrm{e}-07$ | $2.84605 \mathrm{e}-07$ |
| 9 | $1.033829 \mathrm{e}-08$ | $7.834955 \mathrm{e}-09$ | $4.024922 \mathrm{e}-09$ | $1.272792 \mathrm{e}-08$ | $6.363961 \mathrm{e}-09$ | $2.012461 \mathrm{e}-08$ | $2.84605 \mathrm{e}-08$ |
| 10 | $1.033829 \mathrm{e}-09$ | $7.834955 \mathrm{e}-10$ | $4.024922 \mathrm{e}-10$ | $1.272792 \mathrm{e}-09$ | $6.363961 \mathrm{e}-10$ | $2.012461 \mathrm{e}-09$ | $2.84605 \mathrm{e}-09$ |
| 11 | $1.033829 \mathrm{e}-10$ | $7.834955 \mathrm{e}-11$ | $4.024922 \mathrm{e}-11$ | $1.272792 \mathrm{e}-10$ | $6.363961 \mathrm{e}-11$ | $2.012461 \mathrm{e}-10$ | $2.84605 \mathrm{e}-10$ |
| 12 | $1.033829 \mathrm{e}-11$ | $7.834955 \mathrm{e}-12$ | $4.024922 \mathrm{e}-12$ | $1.272792 \mathrm{e}-11$ | $6.363961 \mathrm{e}-12$ | $2.012461 \mathrm{e}-11$ | $2.84605 \mathrm{e}-11$ |
| 13 | $1.033829 \mathrm{e}-12$ | $7.834955 \mathrm{e}-13$ | $4.024922 \mathrm{e}-13$ | $1.272792 \mathrm{e}-12$ | $6.363961 \mathrm{e}-13$ | $2.012461 \mathrm{e}-12$ | $2.84605 \mathrm{e}-12$ |
| 14 | $1.033829 \mathrm{e}-13$ | $7.834955 \mathrm{e}-14$ | $4.024922 \mathrm{e}-14$ | $1.272792 \mathrm{e}-13$ | $6.363961 \mathrm{e}-14$ | $2.012461 \mathrm{e}-13$ | $2.84605 \mathrm{e}-13$ |
| 15 | $1.033829 \mathrm{e}-14$ | $7.834955 \mathrm{e}-15$ | $4.024922 \mathrm{e}-15$ | $1.272792 \mathrm{e}-14$ | $6.363961 \mathrm{e}-15$ | $2.012461 \mathrm{e}-14$ | $2.84605 \mathrm{e}-14$ |

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Table 2 The p.m.f. table of $\mathrm{X} \sim \mathrm{HG}(\mathrm{q}, \theta, \mathrm{k})$, for different k and $\theta$.

| X | $\mathrm{q}=0.1 \theta=0.5$ <br> $\mathrm{k}=10$ | $\mathrm{q}=0.1 \theta=2$ <br> $\mathrm{k}=10$ | $\mathrm{q}=0.25 \theta=0.75$ <br> $\mathrm{k}=5$ | $\mathrm{q}=0.250=0.25$ <br> $\mathrm{k}=5$ | $\mathrm{q}=0.5 \theta=0.1$ <br> $\mathrm{k}=5$ | $\mathrm{q}=0.5 \theta=10$ <br> $\mathrm{k}=5$ |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 0.9066967 | 0.8928227 | 0.7639666 | 0.8105077 | 0.6827162 | 0.2458752 |
| 2 | 0.08397297 | 0.09645961 | 0.1770279 | 0.1421262 | 0.1595167 | 0.3585943 |
| 3 | 0.008397297 | 0.00964596 | 0.0442541 | 0.03552461 | 0.07889698 | 0.1974297 |
| 4 | 0.0008397297 | 0.00096459 | 0.01106353 | 0.00888115 | 0.03943526 | 0.09904511 |
| 5 | $8.397297 \mathrm{e}-05$ | $9.645961 \mathrm{e}-05$ | 0.002765881 | 0.00222028 | 0.01971742 | 0.04952774 |
| 6 | $8.397297 \mathrm{e}-06$ | $9.645961 \mathrm{e}-06$ | 0.0006914703 | 0.000555072 | 0.009858709 | 0.02476395 |
| 7 | $8.397297 \mathrm{e}-07$ | $9.645961 \mathrm{e}-07$ | 0.0001728676 | 0.000138768 | 0.004929354 | 0.01238198 |
| 8 | $8.397297 \mathrm{e}-08$ | $9.645961 \mathrm{e}-08$ | $4.32169 \mathrm{e}-05$ | $3.4692 \mathrm{e}-05$ | 0.002464677 | 0.006190989 |
| 9 | $8.397297 \mathrm{e}-09$ | $9.645961 \mathrm{e}-09$ | $1.080422 \mathrm{e}-05$ | $8.673 \mathrm{e}-06$ | 0.001232339 | 0.003095495 |
| 10 | $8.397297 \mathrm{e}-10$ | $9.645961 \mathrm{e}-10$ | $2.701056 \mathrm{e}-06$ | $2.16825 \mathrm{e}-06$ | 0.0006161693 | 0.001547747 |
| 11 | $8.397297 \mathrm{e}-11$ | $9.645961 \mathrm{e}-11$ | $6.75264 \mathrm{e}-07$ | $5.420625 \mathrm{e}-07$ | 0.0003080846 | 0.0007738736 |
| 12 | $8.397297 \mathrm{e}-12$ | $9.645961 \mathrm{e}-12$ | $1.68816 \mathrm{e}-07$ | $1.355156 \mathrm{e}-07$ | 0.0001540423 | 0.0003869368 |
| 13 | $8.397297 \mathrm{e}-13$ | $9.645961 \mathrm{e}-13$ | $4.2204 \mathrm{e}-08$ | $3.387891 \mathrm{e}-08$ | $7.702116 \mathrm{e}-05$ | 0.0001934684 |
| 14 | $8.397297 \mathrm{e}-14$ | $9.645961 \mathrm{e}-14$ | $1.0551 \mathrm{e}-08$ | $8.469726 \mathrm{e}-09$ | $3.851058 \mathrm{e}-05$ | $9.67342 \mathrm{e}-05$ |
| 15 | $8.397297 \mathrm{e}-15$ | $9.645961 \mathrm{e}-15$ | $2.63775 \mathrm{e}-09$ | $2.117432 \mathrm{e}-09$ | $1.925529 \mathrm{e}-05$ | $4.83671 \mathrm{e}-05$ |

Figure 1 The p.m.f of $\mathrm{X}_{\sim} \mathrm{HG}(\mathrm{q}, \theta, \mathrm{k})$


From (5), Figure 1 and Table $1 \&$ Table 2 it is seen that,
Remark 3 For $\theta<1$ p.m.f. of $\mathrm{X} \sim \mathrm{HG}(\mathrm{q}, \theta, \mathrm{k})$ decreasing with x .

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## III. 3 EXPECTATION, STANDARD DEVIATION AND ENTROPY OF HG DISTRIBUTION

We compute the expectation, standard deviation and Shannon's entropy of the HG distribution with different ' a , k ' and ' $\theta$ ' (Table 3). For a discrete distribution, Shannon's Entropy $=-\sum_{\mathrm{i}} \mathrm{p}\left(\mathrm{x}_{\mathrm{i}}\right) \log \mathrm{p}\left(x_{i}\right), i=1,2 \ldots$

Table 3 Expectation, standard deviation and Shannon's entropy of X~HG (q, $0, \mathrm{k}$ ) (With K=5)

| $\mathrm{k}=5$ | q | $\theta$ | E(x) | Sd(x) | entropy | $\mathrm{k}=5$ | q | $\theta$ | E(x) | Sd(x) | entropy |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 0.9 | 0.75 | 7.181742 | 6.24706 | 2.797387 |  | 0.5 | 2 | 2.145117 | 1.459823 | 1.476805 |
|  | 0.75 | 0.75 | 3.822205 | 3.363209 | 2.194411 |  | 0.1 | 2 | 1.127633 | 0.3737722 | 0.3980682 |
|  | 0.1 | 0.75 | 1.104899 | 0.3423532 | 0.346722 |  | 0.9 | 5 | 9.231612 | 6.478768 | 2.88427 |
|  | 0.9 | 0.5 | 6.761051 | 6.181877 | 2.741898 |  | 0.75 | 5 | 4.9415 | 3.514446 | 2.425729 |
|  | 0.75 | 0.5 | 3.61276 | 3.310981 | 2.120546 |  | 0.1 | 5 | 1.153302 | 0.4048062 | 0.4511009 |
|  | 0.1 | 0.5 | 1.096728 | 0.3299496 | 0.3271156 |  | 0.9 | 10 | 9.954691 | 6.57884 | 2.858474 |
|  | 0.75 | 0.2 | 3.186474 | 3.169406 | 1.934354 |  | 0.75 | 10 | 5.388573 | 3.538493 | 2.465914 |
|  | 0.1 | 0.2 | 1.080531 | 0.3032189 | 0.2861816 |  | 0.5 | 10 | 2.545866 | 1.524628 | 1.628365 |
|  | 0.9 | 2 | 8.238283 | 6.372622 | 2.875304 |  | 0.1 | 10 | 1.176096 | 0.4292081 | 0.4943981 |
|  | 0.75 | 2 | 4.376501 | 3.459642 | 2.338958 |  | 0.9 | 100 | 11.89503 | 7.275485 | 2.633785 |

A closer look with the entropy of $\mathrm{X} \sim \mathrm{HG}(\mathrm{q}, \theta, \mathrm{k})$ with different $\mathrm{q}, \theta$ and k (table 4 ).
Table 4 Entropy of $\mathrm{X} \sim \mathrm{HG}(\mathrm{q}, \theta, \mathrm{k})$ with different $\mathrm{q}, \mathrm{k} \& \theta$, $(\mathrm{k}=2)$.

| k | q | $\theta=0.2$ | $\theta=0.5$ | $\theta=2$ | $\theta=5$ | $\theta=10$ | k | q | $\theta=0.2$ | $\theta=0.5$ | $\theta=2$ | $\theta=5$ | $\theta=10$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2 | 0.9 | 2.549 | 2.988 | 3.450 | 3.631 | 3.718 | 5 | 0.9 | 2.891 | 3.118 | 3.349 | 3.435 | 3.475 |
| 2 | 0.75 | 1.608 | 1.997 | 2.447 | 2.629 | 2.717 |  | 0.75 | 1.939 | 2.126 | 2.347 | 2.435 | 2.476 |
|  | 0.5 | 0.892 | 1.171 | 1.575 | 1.758 | 1.848 |  | 0.5 | 1.159 | 1.289 | 1.476 | 1.576 | 1.628 |
|  | 0.1 | 0.199 | 0.281 | 0.457 | 0.606 | 0.726 |  | 0.1 | 0.286 | 0.327 | 0.398 | 0.451 | 0.494 |

Remark 4 When $k$ increases, the entropy of $X \sim H G(q, \theta, k)$ increases with a fixed $\theta$ and $q$ when $(\theta<1)$ and decreases with a fixed $\theta$ and $q$ when $(\theta>1)$. When $\theta$ increases the entropy of $X \sim H G(q, \theta, k)$ increases with a fixed $q$ and $k$. Also when $q$ increases the entropy of $\mathrm{X} \sim \mathrm{HG}(\mathrm{q}, \theta, \mathrm{k})$ decreases with a fixed $\theta$ and k .

Remark 5 For $X \sim H G(q, \theta, k)$ the $E(X)$, the $\operatorname{sd}(X)$ and the entropy are greater than the $E(X)$,the $\operatorname{sd}(X)$ and the entropy of $X \sim \operatorname{MOG}(\mathrm{q}, \theta)$ when $\theta<1$ and less than the $\mathrm{E}(\mathrm{X})$,the $\operatorname{sd}(\mathrm{X})$ and the entropy of $\mathrm{X} \sim \operatorname{MOG}(\mathrm{q}, \theta)$ when $\theta>1$.

Mean, median and mode of the HG distribution: Mean, median and mode of the HG distribution with different $q$ and $\theta(\mathrm{n}=100)$ are calculated (Table 5) from Table $1 \&$ Table 2 and Figure 1.

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Table 5 Mean, median and mode of the HG distribution with different $q$ and $\theta(\mathrm{n}=100)$

| k | q | $\theta$ | Mean | Median | Mode | k | q | $\theta$ | Mean | Median | Mode |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 5 | 0.1 | 0.1 | 1.07 | 1 | 1 | 5 | 0.75 | 0.1 | 2.92 | 1 | 1 |
|  |  | 0.25 | 1.084 | 1 | 1 |  |  | 0.25 | 3.3 | 2 | 1 |
|  |  | 0.5 | 1.09 | 1 | 1 |  |  | 0.5 | 3.63 | 2 | 1 |
|  |  | 2 | 1.081 | 1 | 1 |  |  | 2 | 4.401 | 3 | 2 |
|  |  | 4 | 1.092 | 1 | 1 |  |  | 4 | 4.829 | 4 | 2 |
|  |  | 10 | 1.176 | 1 | 1 |  |  | 10 | 5.423 | 4 | 3 |
|  | 0.5 | 0.1 | 1.632 | 1 | 1 |  | 0.9 | 0.1 | 6.849 | 3 | 1 |
|  |  | 0.25 | 1.759 | 1 | 1 |  |  | 0.25 | 7.98 | 5 | 1 |
|  |  | 0.5 | 1.871 | 1 | 1 |  |  | 0.5 | 8.945 | 6 | 1 |
|  |  | 2 | 2.145 | 2 | 1 |  |  | 2 | 11.12 | 8 | 4 |
|  |  | 4 | 2.307 | 2 | 1 |  |  | 4 | 12.3 | 10 | 6 |
|  |  | 10 | 2.545 | 2 | 2 |  |  | 10 | 13.45 | 11 | 8 |

From table 5 we observed that for $\mathrm{X} \sim \mathrm{HG}(\mathrm{a}, \theta, \mathrm{k})$ the mean is always greater than median and mode, hence, Remark 6 For $\mathrm{X} \sim \mathrm{HG}(\mathrm{a}, \theta, \mathrm{k})$ the mean is greater than median and mode, hence the HG distribution is always (irrespective of the value of $\theta$ )positively skewed. This suggests that the HG distribution can be a good fit for positively skewed data.
Remark 7 For $\mathrm{X} \sim \mathrm{HG}(\mathrm{q}, \theta, \mathrm{k})$ the mean, median and mode are less than or equal to the mean, median and mode of $\mathrm{X} \sim$ $\operatorname{MOG}(\mathrm{q}, \theta) .($ here $\mathrm{k}=1)$

## IV. 4 HAZARD FUNCTION

The hazard function, $\gamma_{H}(x)=h(x) / H(x)=[(H(x-1, \theta, k)) /(H(x, \theta, k))]-1$

$$
\begin{align*}
& \left.\left.=\left[\theta^{1 / k} q^{\mathrm{x}-1} /\left[1-(1-\theta) q^{(\mathrm{x}-1) \mathrm{k}}\right]^{1 / \mathrm{k}}\right]\right] /\left[\theta^{1 / \mathrm{k}} \mathrm{q}^{\mathrm{x}} /\left[1-(1-\theta) \mathrm{q}^{\mathrm{xk}}\right]^{1 / \mathrm{k}}\right]\right]-1 \\
& =(1 / \mathrm{q})\left\{\left[1-(1-\theta) \mathrm{q}^{\mathrm{xk}}\right] /\left[1-(1-\theta) \mathrm{q}^{(\mathrm{x}-1) \mathrm{k}}\right]\right\}^{1 / \mathrm{k}}-1 . \tag{6}
\end{align*}
$$

The hazard function for different $p, k \& \theta$ is calculated and represented in Table 7. The graph of the Hazard function of X , with different values for the parameters $\mathrm{q}, \theta \& \mathrm{k}$ is shown in Figure 2.

Table 7Hazard function $\gamma_{\mathrm{H}}(\mathrm{x})$ of $\mathrm{X} \sim \mathrm{HG}(\mathrm{q}, \theta, \mathrm{k}), \mathrm{q}=0.75, \theta<1$

| $\mathrm{q}=0.75 \mathrm{k}=5$ |  |  |  |  |  | $\mathrm{q}=0.75 \mathrm{k}=5$ |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| x | $\theta=0.75$ | $\theta=0.5$ | $\theta=0.25$ | $\theta=0.05$ | x | $\theta=2$ | $\theta=5$ | $\theta=10$ | $\theta=50$ |  |
| 1 | 0.395128 | 0.493393 | 0.691715 | 1.306512 | 1 | 0.211234 | 0.104375 | 0.057320 | 0.012537 |  |
| 2 | 0.345920 | 0.359654 | 0.374718 | 0.387880 | 2 | 0.291819 | 0.215099 | 0.151550 | 0.046393 |  |
| 3 | 0.336225 | 0.339174 | 0.342183 | 0.344635 | 3 | 0.322309 | 0.293649 | 0.256582 | 0.131528 |  |
| 4 | 0.334014 | 0.334698 | 0.335386 | 0.335938 | 4 | 0.330640 | 0.322849 | 0.310745 | 0.240894 |  |
| 5 | 0.333494 | 0.333656 | 0.333817 | 0.333947 | 5 | 0.332689 | 0.330775 | 0.327641 | 0.304783 |  |
| 6 | 0.333371 | 0.333409 | 0.333448 | 0.333478 | 6 | 0.333180 | 0.332722 | 0.331962 | 0.326020 |  |
| 7 | 0.333342 | 0.333351 | 0.333360 | 0.333367 | 7 | 0.333297 | 0.333188 | 0.333006 | 0.331564 |  |
| 8 | 0.333335 | 0.333337 | 0.333339 | 0.333341 | 8 | 0.333324 | 0.333298 | 0.333255 | 0.332911 |  |
| 9 | 0.333333 | 0.333334 | 0.333334 | 0.333335 | 9 | 0.333331 | 0.333325 | 0.333314 | 0.333233 |  |
| 10 | 0.333333 | 0.333333 | 0.333333 | 0.333333 | 10 | 0.333332 | 0.333331 | 0.333329 | 0.333309 |  |
| 11 | 0.333333 | 0.333333 | 0.333333 | 0.333333 | 11 | 0.333333 | 0.333332 | 0.333332 | 0.333327 |  |
| 12 | 0.333333 | 0.3333333 | 0.333333 | 0.333333 | 12 | 0.333333 | 0.333333 | 0.333333 | 0.333332 |  |

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Figure 2 Hazard function of $\mathrm{X} \sim \mathrm{HG}(\mathrm{q}, \theta, \mathrm{k})$


It shows when $\theta>1$ the failure rate is non-decreasing and when $\theta<1$ it is non-increasing. When $x$ tends to 1 , it is seen that the hazard rate tends to a constant equal to( $1 / \mathrm{q})-1=(\mathrm{p} / \mathrm{q})$ for any $\theta>0$.
From (6), Table 7 and Figure 2,
Remark 8 For $\mathrm{X} \sim \mathrm{HG}(\mathrm{q}, \theta, \mathrm{k})$, when q is small the hazard rate tends to a constant equal to $(1 / \mathrm{q})-1=(\mathrm{p} / \mathrm{q})$. i.e., $\gamma_{\mathrm{H}}(\mathrm{x})$ tends to $\mathrm{p} / \mathrm{q}$, irrespective of the value of $\theta$.

Proof: For $X \sim H G(q, \theta, k)$, from $(6), \gamma_{H}(x)=(1 / q)\left\{\left[1-(1-\theta) q^{x k}\right] /\left[1-(1-\theta) q^{(x-1) k}\right]\right\}^{1 / k}-1$.
i.e., $\gamma_{H}(x)=(1 / q)\left\{\left[1-(1-\theta) q^{x k}\right] /\left[1-(1-\theta) q^{x k} / q^{k}\right]\right\}^{1 / k}-1=(1 / q)\left\{\left[1-(1-\theta) q^{x k}\right] /\left[\left(q^{k}-(1-\theta) q^{\mathrm{xk}}\right) / q^{k}\right]\right\}^{1 / k}-1$.
i.e., as $x$ tends to $\infty, \gamma_{H}(x)$ tends to $(1 / q)-1$ (here ( $1 / \mathrm{q}$ ) $-1=\mathrm{p} / \mathrm{q}$ )

From the simulation results, it is seen that the rate of convergence is more rapidly when k is large.
Increasing /Decreasing failure rate (IFR/DFR): Let us define $\eta(t)=1-[p(x=t+1) / p(x=t)] \& \Delta \eta(t)=\eta(t+1)-\eta(t)$. Kemp[4] showed, If $\Delta \eta(t)>0$, the distribution is with IFR, if $\Delta \eta(t)<0$ the distribution is with DFR and if $\Delta \eta(t)=0$ the distribution is with constant FR. Also $\Delta \eta(t)=\eta(t+1)-\eta(t)=[P(x=t+1) / P(x=t)]-[P(x=t+2) / P(x=t+1)]>0$, when $\mathrm{P}(\mathrm{x}=\mathrm{t}+1)^{2}>\mathrm{P}(\mathrm{x}=\mathrm{t}) \mathrm{P}(\mathrm{x}=\mathrm{t}+2)$, then, the probabilities are log-concave and if $[\mathrm{P}(\mathrm{X}=\mathrm{x}+1)]^{2}<\mathrm{P}(\mathrm{X}=\mathrm{x}) \mathrm{P}(\mathrm{X}=\mathrm{x}+2)$ then probabilities are log-convex. By definition 2.1 in (Sandhya and Prasanth [8])

Here, $1-\eta(\mathrm{t})=\mathrm{h}(\mathrm{t}+1, \theta, \mathrm{k}) / \mathrm{h}(\mathrm{t}, \theta, \mathrm{k})$

$$
=\theta^{1 / k} q^{t}\left\{\left[1 /\left(1-(1-\theta) q^{\mathrm{tk}}\right)^{1 / k}\right]-\left[q /\left(1-(1-\theta) q^{(t+1) k}\right)^{1 / k}\right]\right\} / \theta^{1 / k} q^{t-1}\left\{\left[1 /\left(1-(1-\theta) q^{(t-1) k}\right)^{1 / k}\right]-\left[q /\left(1-(1-\theta) q^{\mathrm{tk}}\right)^{1 / k}\right]\right\} .
$$

$1-\eta(t+1)=h(t+2, \theta, k) / h(t+1, \theta, k)$

$$
=\theta^{1 / k} q^{t+1}\left\{\left[1 /\left(1-(1-\theta) q^{(t+1) k}\right)^{1 / k}\right]-\left[q /\left(1-(1-\theta) q^{(t+2) k}\right)^{1 / k}\right]\right\} / \theta^{1 / k} q^{t}\left\{\left[1 /\left(1-(1-\theta) q^{t k}\right)^{1 / k}\right]-\left[q /\left(1-(1-\theta) q^{(t+1) k}\right)^{1 / k}\right]\right\}
$$

$\Delta \eta(\mathrm{t})=\eta(\mathrm{t}+1)-\eta(\mathrm{t})=\{1-[\mathrm{h}(\mathrm{t}+2, \theta, \mathrm{k}) / \mathrm{h}(\mathrm{t}+1, \theta, \mathrm{k})]\}-\{1-[\mathrm{h}(\mathrm{t}+1, \theta, \mathrm{k}) / \mathrm{h}(\mathrm{t}, \theta, \mathrm{k})]\}$
That is when $\mathrm{X} \sim \mathrm{HG}(\mathrm{q}, \theta, \mathrm{k}), \Delta \eta(\mathrm{t})=\{\mathrm{h}(\mathrm{t}+1, \theta, \mathrm{k}) / \mathrm{h}(\mathrm{t}, \theta, \mathrm{k})\}-\{\mathrm{h}(\mathrm{t}+2, \theta, \mathrm{k}) / \mathrm{h}(\mathrm{t}+1, \theta, \mathrm{k})\}$.
We can check it with different $\theta, \mathrm{k}$ and q . then calculated $\Delta \eta(\mathrm{t})$ in Table 8 and Table 9.

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Table 8 Calculated $\Delta \eta(t)$ of $X \rightarrow H G(q ; \theta ; k), \theta<1$

|  | $\mathrm{q}=0.9 \mathrm{k}=5$ |  | $\mathrm{q}=0.9 \mathrm{k}=25$ |  |
| :--- | :---: | :---: | :---: | :---: |
| x | $\theta=0.75$ | $\theta=0.05$ | $\theta=0.75$ | $\theta=0.05$ |
| 1 | -0.0429893 | -0.3943933 | -0.07748027 | -0.4672989 |
| 2 | -0.02335815 | -0.1290236 | -0.005229174 | -0.0201156 |
| 3 | -0.01315994 | -0.06107629 | -0.0003739249 | -0.001422121 |
| 4 | -0.007563441 | -0.03216002 | $-2.683646 \mathrm{e}-05$ | -0.0001019847 |
| 5 | -0.004396362 | -0.01782131 | $-1.926545 \mathrm{e}-06$ | $-7.320903 \mathrm{e}-06$ |
| 6 | -0.002572178 | -0.01014861 | $-1.383061 \mathrm{e}-07$ | $-5.255633 \mathrm{e}-07$ |
| 7 | -0.001510637 | -0.00586810 | $-9.928963 \mathrm{e}-09$ | $-3.773006 \mathrm{e}-08$ |
| 8 | -0.0008891741 | -0.00342281 | $-7.127966 \mathrm{e}-10$ | $-2.708635 \mathrm{e}-09$ |
| 9 | -0.0005240617 | -0.00200664 | $-5.117307 \mathrm{e}-11$ | $-1.944526 \mathrm{e}-10$ |
| 10 | -0.00030911 | -0.00117989 | $-3.674283 \mathrm{e}-12$ | $-1.395883 \mathrm{e}-11$ |
| 11 | -0.0001824069 | -0.00069498 | $-2.63678 \mathrm{e}-13$ | $-1.003309 \mathrm{e}-12$ |
| 12 | -0.0001076678 | -0.00040977 | $-1.798561 \mathrm{e}-14$ | $-7.083223 \mathrm{e}-14$ |

Table 9 Calculated $\Delta \eta(\mathrm{t})$ of $\mathrm{X} \sim \mathrm{HG}(\mathrm{q}, \theta, \mathrm{k}), \theta>1$.

|  | $\mathrm{q}=0.9 \mathrm{k}=5$ |  | $\mathrm{q}=0.9 \mathrm{k}=25$ |  |
| :--- | :---: | :---: | :---: | :---: |
| x | $\theta=5$ | $\theta=50$ | $\theta=5$ | $\theta=50$ |
| 1 | 0.1012197 | 0.02357858 | 0.7631403 | 4.02686 |
| 2 | 0.1012504 | 0.0366706 | 0.07843443 | 0.6907295 |
| 3 | 0.08900311 | 0.05422491 | 0.005952379 | 0.06943551 |
| 4 | 0.06975785 | 0.07451798 | 0.000429225 | 0.005237675 |
| 5 | 0.04987859 | 0.09279671 | $3.08239 \mathrm{e}-05$ | 0.000377487 |
| 6 | 0.03334167 | 0.1024701 | $2.212893 \mathrm{e}-06$ | $2.710739 \mathrm{e}-05$ |
| 7 | 0.02128242 | 0.0992576 | $1.588634 \mathrm{e}-07$ | $1.946074 \mathrm{e}-06$ |
| 8 | 0.01318267 | 0.08469318 | $1.140477 \mathrm{e}-08$ | $1.397085 \mathrm{e}-07$ |
| 9 | 0.008012775 | 0.06476242 | $8.187458 \mathrm{e}-10$ | $1.002964 \mathrm{e}-08$ |
| 10 | 0.004814235 | 0.04545278 | $5.877621 \mathrm{e}-11$ | $7.20028 \mathrm{e}-10$ |
| 11 | 0.002872283 | 0.02999003 | $4.2224 \mathrm{e}-12$ | $5.168732 \mathrm{e}-11$ |
| 12 | 0.00170649 | 0.01897888 | $3.009815 \mathrm{e}-13$ | $3.714362 \mathrm{e}-12$ |

So it's clear that, for $\theta>1, \Delta \eta(t)>0$, and implies then the distribution is with IFR. Similarly when $\theta<1, \Delta \eta(t)<0$, the distribution is with DFR and when $\theta=1$, then $\Delta \eta(\mathrm{t})=0$,the distribution is with constant FR. This constant FR is more often when $X$ is large and when $q$ is small the hazard rate tends to a constant equal to $(1 / \mathrm{q})-1=(\mathrm{p} / \mathrm{q})$, (remark8) i.e. $\gamma_{\mathrm{H}}(\mathrm{x})$ tends to $\mathrm{p} / \mathrm{q}$. This is more rapidly when k is large.

Hence $[P(X=x+1)]^{2}<P(X=x) P(X=x+2)$ when $\theta<1$.
i.e., probabilities are log-convex when $\theta<1$ and it is also seen that,
$[P(X=x+1)]^{2}>P(X=x) P(X=x+2)$ when $\theta>1$.
i.e., probabilities are log-concave when $\theta>1$

Remark 9 In discrete case, following Steutel [14], for distributions on non negative integers, log-convexity implies that they are compound geometric and hence they are infinitely divisible(i.d.) From Klebanov ,et.al [5], it is clear that compound geometric distributions are Geometric infinitely divisible(GID) and that log-convexity implies that the distribution has decreasing hazard rate(DHR) (Johnson et. al [3], p.209).

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Thus we have the following proposition
Proposition1 If a discrete random variable $X_{\sim} \operatorname{HG}(q, \theta), x=1,2,3, \ldots$ then, when $0<\theta<1$, the distribution is i) log-convex ii)compound geometric ,iii)infinitely divisible (i.d.), iv)geometrically infinite divisible (GID), and with v) decreasing hazard rate (DHR).

Proof of (i) is clear from (7), and the remaining follows from the remark 9 and is similar to that of Theorem2.1, Sandhya and Prasanth [8].

New better than used/New worse than used (NBU/NWU): We know that, if the reliability function $H_{t+x}<H_{t} H_{x}$ then NBU and if $\mathrm{H}_{\mathrm{t}+\mathrm{x}}>\mathrm{H}_{\mathrm{t}} \mathrm{H}_{\mathrm{x}}$ then NWU. (Adrinue W.Kemp [4]) explained the characteristics of discrete life-time distributions)

Let $X \sim H(q, e, k)$ then, $H_{x}, H_{t}$ and $H_{t+x}$ are the s.fs of $X=x, X=t$ and $X=x+t$ respectively.

$$
\mathrm{H}_{t+\mathrm{x}}=\theta^{1 / k} \mathrm{q}^{\mathrm{t}+\mathrm{x}} /\left[1-(1-\theta) \mathrm{q}^{(\mathrm{t+x}) \mathrm{k}}\right]^{1 / k}, \mathrm{H}_{\mathrm{x}}=\theta^{1 / k} \mathrm{q}^{\mathrm{x}} /\left[1-(1-\theta) \mathrm{q}^{\mathrm{xk}}\right]^{1 / \mathrm{k}}, \mathrm{H}_{\mathrm{t}}=\theta^{1 / \mathrm{k}} \mathrm{q}^{\mathrm{t} /\left[1-(1-\theta) \mathrm{q}^{\mathrm{tk}}\right]^{1 / k} .}
$$

If $\theta>1$, then, $\mathrm{H}_{\mathrm{t}+\mathrm{x}}<\mathrm{H}_{\mathrm{t}} \mathrm{H}_{\mathrm{x}}$ then the distribution is NBU and if $\theta<1$, the distribution is NWU. Proof is similar to that of MOG distribution (Sandhya and Prasanth [8]).

Remark 10 For the Harris-Geometric distribution,

```
when 0>1, the distribution is log-convex }->\mathrm{ IFR }->\mathrm{ IFRA }->\mathrm{ NBU }->\mathrm{ NBUE }->\mathrm{ DMRL &
when }0<1\mathrm{ , the distribution is log-concave }->\mathrm{ DFR }->\mathrm{ DFRA }->\mathrm{ NWU }->\mathrm{ NWUE }->\mathrm{ IMRL. (By Remark 9)
```

Remark 11 Let $\gamma_{\mathrm{F}(\mathrm{x})}$ and $\gamma_{\mathrm{H}(\mathrm{x}, \theta)}$ are the Hazard rate of the geometric distribution and HG distribution respectively. Then, i) $\operatorname{Limit}_{\left(\mathrm{x}_{\rightarrow}^{\alpha}\right) \gamma \mathrm{H}(\mathrm{x}, \theta)}=\operatorname{Limit}_{\left(\mathrm{x}_{\rightarrow^{\alpha}}\right)} \gamma_{\mathrm{F}(\mathrm{x}, \theta)}$ for any $\theta \geq 0$. ii) $\gamma_{\mathrm{F}(\mathrm{x})} / \theta \leq \gamma_{\mathrm{H}(\mathrm{x}, \theta)} \leq \gamma_{\mathrm{F}(\mathrm{x})}, \theta \geq 1$. iii) $\gamma_{\mathrm{F}(\mathrm{x})} \leq \gamma_{\mathrm{H}(\mathrm{x}, \theta)} \leq \gamma_{\mathrm{F}(\mathrm{x})} / \theta, \quad 0$ $\leq \theta \leq 1$. (Proof is so direct from 1.2 and 4.1).

## V. 5. HG DISTRIBUTION AS THE DISTRIBUTION OF MINIMUM OF A SEQUENCE OF I.I.D. VARIABLES

The following theorem gives characterization of HG distribution as the minimum of a sequence of i.i.d. variables following geometric distribution.

Theorem 1 Let $\left\{X_{i}, i \geq 1\right\}$ be a sequence of i.i.d. r.vs with common s.f $\mathrm{F}(\mathrm{x})$. Let N be a Harris r.v $\mathrm{H} 1(\theta, \mathrm{k} ; 1 / \mathrm{k})$ independent of $\left\{\mathrm{X}_{\mathrm{i}}, \mathrm{i} \geq 1\right\}$. such that $\mathrm{P}(\mathrm{N}=\mathrm{n})=\left[{ }^{(\mathrm{n} / \mathrm{k})-1} \mathrm{C}_{(\mathrm{n} / \mathrm{k})-1}\right](\theta)^{1 / k}(1-\theta)^{(\mathrm{n}-1) / \mathrm{k})}, \mathrm{n}=1,1+\mathrm{k}, 1+2 \mathrm{k}, \ldots \mathrm{k}>0$ integer, $0<\theta<1$. Let $\mathrm{U}_{\mathrm{N}}=\min _{1 \leq i \leq n}\left(\mathrm{X}_{\mathrm{i}}\right)$. Then $\left\{\mathrm{U}_{\mathrm{N}}\right\}$ is distributed as Harris Geometric distribution if and only if $\left\{\mathrm{X}_{\mathrm{i}}\right\} \sim$ geometric $(p)$.

Proof: Let the s.f of $\mathrm{U}_{\mathrm{N}}$ is,
$J(x)=P\left(U_{N}>x\right)=\sum_{1}{ }^{\infty}[F(x)]^{n} P(X=x)=\theta^{1 / k} F(x) /\left[1-(1-\theta) F^{k}(x)\right]^{1 / k}=\theta^{1 / k}\left[\left(q^{x}\right] /\left[1-(1-\theta) q^{x k}\right]^{1 / k}=\theta^{1 / k} /\left[q^{-k x}-(1-\theta)\right]^{1 / k}\right.$.
Which is the s.f of Harris- Geometric ( $\mathrm{q}, \theta, \mathrm{k}$ ).Retracing the steps we get the only if part.

## VI. 6 AR (1) MODEL WITH HG DISTRIBUTION AS INNOVATION DISTRIBUTION

Consider k independent $\operatorname{AR}(1)$ sequences $\{X i, n\}, \mathrm{i}=1,2, \ldots \mathrm{k}$, for all $\mathrm{n}>0$ integer and some $\mathrm{b}>0$
$\Lambda_{\mathrm{i}} \mathrm{X}_{\mathrm{i}, \mathrm{n}}=\mathrm{b}\left\{\Lambda_{\mathrm{i}} \mathrm{X}_{\mathrm{i}, \mathrm{n}-1}\right\}$ with probability $p$

$$
\begin{equation*}
=\mathrm{b}\left\{\Lambda_{\mathrm{i}} \mathrm{X}_{\mathrm{i}, \mathrm{n}-1}\right\} \Lambda\left\{\Lambda_{\mathrm{i}} \varepsilon_{\mathrm{i}, \mathrm{n}-1}\right\} \text { with probability 1-p } \tag{9}
\end{equation*}
$$

The structure (9) was defined by Satheesh et al [11].

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Theorem 2 In an AR (1) process with structure (9), $\left\{\mathrm{X}_{\mathrm{i}, \mathrm{n}}\right\}$ is stationary Markovian with Harris Geometric ( $\theta, \mathrm{a}, \mathrm{k}$ ) marginal if and only if $\left\{\Theta_{\mathrm{n}}\right\}$ is distributed as Geometric distribution with parameter $(q)$.

Proof $\mathrm{H}(\mathrm{t})=\left\{\theta \mathrm{F}^{\mathrm{k}}(\mathrm{bt}) /\left[1-(1-\theta) \mathrm{F}^{\mathrm{k}}(\mathrm{bt})\right\}^{1 / \mathrm{k}}\right.$. Then let $\mathrm{F}^{\mathrm{k}}(\mathrm{bt})=\mathrm{q}^{\mathrm{x}}$, then
$H(t)=\left\{\theta F^{k}(b t) /\left[1-(1-\theta) F^{k}(b t)\right\}^{1 / k}=\left[\theta^{1 / k} q^{k x} /\left[1-(1-\theta) q^{k x}\right]^{1 / k}=\theta^{1 / k} /\left[q^{-k x}-(1-\theta)\right]^{1 / k}\right.\right.$. So $X_{i, n}$ is $H G(q, \theta, k)$. We can-show the converse by retracing the steps.

## VII. 7 MAXIMUM LIKELIHOOD ESTIMATES (M.L.E.) OF THE PARAMETERS OF HG (Q,@,K)

Let $\mathrm{x}_{1}, \mathrm{x}_{2} \ldots \mathrm{x}_{\mathrm{n}} \mathrm{n}$ independent random samples from $\mathrm{HG}(\mathrm{q}, \theta, \mathrm{k})$. Then from (5) we write the likelihood function L of the distribution. By partial differentiation of L w.r.to $\mathrm{q}, \theta \& \mathrm{k}$ and equate them to zero we formulate 3 non-linear equations. We calculated the m.l.e. of $\mathrm{q}, \theta$ and k numerically (using Mathematica) as the solution of these three nonlinear equations.

## VIII. 8 SOME APPLICATIONS OF HG DISTRIBUTION

Example 1.Sales of medicine - domcet / dompar/cetadom (domperidone(10mg) with paracetamol ( 500 mg )) in each 50days (excluding Sundays /each 2 months of one year ) from a medical shop in the Palghat city(in Kerala, India) are noted. Here $\mathrm{X}=$ number of customers (with or without medical prescription brought the medicine) for domcet/dompar /cetadom medicine (domperidone with paracetamol).

Table 10 Observed frequencies in each 50 days ( 2 months)

| x |  | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 | 13 | 14 | 15 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| month | May-June | 25 | 8 | 8 | 1 | 2 | 0 | 4 | 1 | 0 | 1 | 0 | 0 | 0 | 0 | 0 |
|  | July- August | 21 | 7 | 6 | 4 | 2 | 1 | 2 | 3 | 1 | 0 | 2 | 0 | 0 | 0 | 0 |
|  | September- October | 26 | 9 | 4 | 3 | 3 | 0 | 1 | 4 | 0 | 1 | 0 | 0 | 0 | 0 | 0 |
|  | November- December | 24 | 9 | 3 | 2 | 5 | 2 | 1 | 1 | 2 | 0 | 0 | 1 | 0 | 0 | 0 |
|  | January- February | 26 | 9 | 3 | 3 | 2 | 2 | 1 | 1 | 2 | 0 | 0 | 0 | 0 | 0 | 0 |
|  | March - April | 24 | 10 | 1 | 5 | 2 | 2 | 2 | 0 | 1 | 0 | 1 | 1 | 0 | 1 | 0 |

We have analyzed all these data sets. For convenience we arbitrarily take 3 month groups and represent the results here. We observe the mean, median and mode of these data (Table 11).

Table 11 Mean median and mode of the data in May-June, Nov-Dec and March-April

| Month | mean | median | mode |
| :---: | :---: | :---: | :---: |
| May-June | 2.48 | 1 | 1 |
| Nov-Dec | 2.8 | 2 | 1 |
| March-April | 2.98 | 2 | 1 |

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Since mean > mode and mean > median, the data exhibits positive skeweness and it is observed that the distribution is unimodal, hence HG distribution is supposed to give a better fit.

Table 12Goodness of fit of HG distribution using chi-square test with $\alpha=0.05$

|  | May-June | Nov-Dec | March-April |
| :---: | :---: | :---: | :---: |
| m.l.e q^ | 0.7501 | 0.7256 | 0.7697 |
| m.l.e. $\theta^{\wedge}$ | 0.1888 | 0.2250 | 0.2151 |
| m.l.ek | $3(3.1106)$ | $3(3.1695)$ | $3(2.5632)$ |
| Observed $\chi^{\wedge}$ | 4.294 | 2.0125 | 3.91 |
| degrees freedom d.f. | 4 | 4 | 4 |
| p-value | 0.3677 | 0.7335 | 0.4183 |

Result: From the p-values it is seen that here HG distribution is suitable here.
Example 2. Sales of medicine for kids - lanol / calpol / fepanil/ paracetamo/ flucold (with paracetamol ( $60 \mathrm{ml}-125 \mathrm{mg}$ )) within each 50days (excluding Sundays in each 2months of one year) from a medical shop in the Palghat city are noted .Here $\mathrm{X}=$ number of customers (with or without medical prescription brought the medicine) for lanol / calpol / fepanil/ paracetamol/ flucold (with paracetamol ( $60 \mathrm{ml}-125 \mathrm{mg}$ )).

Table 13 Observed frequencies in each 50 days (2 months)

| x | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 | 13 | 14 | 15 | 16 | 17 | 18 | 19 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Month |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| May-June | 12 | 10 | 3 | 4 | 2 | 4 | 2 | 1 | 3 | 2 | 2 | 1 | 1 | 0 | 1 | 0 | 0 | 0 | 2 |
| July- August | 10 | 11 | 3 | 0 | 5 | 4 | 3 | 0 | 1 | 2 | 4 | 0 | 1 | 2 | 1 | 1 | 0 | 1 | 1 |
| September- October | 11 | 8 | 6 | 3 | 2 | 3 | 2 | 2 | 3 | 2 | 1 | 0 | 3 | 3 | 1 | 0 | 0 | 0 | 0 |
| November- December | 12 | 6 | 8 | 1 | 3 | 2 | 4 | 1 | 2 | 1 | 2 | 1 | 1 | 1 | 1 | 0 | 1 | 2 | 1 |
| January- February | 8 | 11 | 1 | 4 | 4 | 2 | 2 | 4 | 3 | 2 | 0 | 2 | 0 | 1 | 2 | 2 | 0 | 0 | 0 |
| March - April | 14 | 6 | 10 | 5 | 2 | 3 | 5 | 0 | 1 | 0 | 2 | 0 | 1 | 0 | 0 | 1 | 0 | 0 | 0 |

We analyzed all these data sets. For convenience we arbitrarily take 3 month groups and represent the results here. We observe the mean, median and mode of these data (Table 14).

Table 14 Mean median and mode of the data in May-June, September- October and March-April

| Month | Mean | median | mode |
| :---: | :---: | :---: | :---: |
| May-June | 5.2 | 3 | 1 |
| Sep- Oct | 5.38 | 3 | 1 |
| March-April | 4.8 | 3 | 1 |

Since mean > mode and mean > median, the data exhibits positive skeweness with m.l.e. of $\theta>0$ and it is observed that the distribution is unimodal, hence HG distribution is supposed to give a better fit.

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Table 15 Goodness of fit of HG distribution using chi-square test with $\alpha=0.05$

|  | May-June | Sep- Oct | March-April |
| :---: | :---: | :---: | :---: |
| MLE q^ | 0.8474 | 0.8366 | 0.8327 |
| MLE $\theta^{\wedge}$ | 0.5142 | 0.5784 | 0.4685 |
| MLE $\mathrm{k}^{\wedge}$ | $3.5824(4)$ | $4.2426(4)$ | $4.1448(4)$ |
| Observed $\chi^{2}$ | 2.6766 | 3.492 | 9.74357 |
| d.f. | 3 | 4 | 6 |
| p-value | 0.444 | 0.4791 | 0.136 |

Result: From the p-values it is seen that here HG distribution is suitable
Example3 Sales of medicine for kids - lanol / calpol / fepanil/ paracetamol/ flucold (with paracetamol ( $60 \mathrm{ml}-125 \mathrm{mg}$ )) with any of the anti-biotic like Moxikind dispertab (amoxicillin trihydrate 125 mg ) or any (with Azithromycin/Erythromycin/ Cefoperazonie/ Ampicllin/ Amoxicillin) within each 50days (excluding Sundays /each 2 months of one year ) from a medical shop in the Palghat city are taken. Here $\mathrm{X}=$ number of customers (with or without medical prescription brought the medicine) for lanol / calpol / fepanil/ paracetamol/ flucold (with paracetamol ( $60 \mathrm{ml}-125 \mathrm{mg}$ )) with any of anti-biotic like Moxikind dispertab(amoxicillin trihydrate 125 mg ) or (Azithromycin/ Erythromycin/Cefoperazone/Ampicillin/ Amoxicillin) within each 50 days

Table 16 Observed frequencies in each 50 days (2 months)

| month | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 | 13 | 14 | 15 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| May- June | 13 | 19 | 6 | 7 | 5 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| July- August | 11 | 15 | 13 | 5 | 3 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| September- October | 12 | 16 | 13 | 4 | 2 | 1 | 1 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 |
| November- December | 9 | 18 | 11 | 6 | 3 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 2 | 0 | 0 |
| January- February | 8 | 12 | 10 | 13 | 3 | 1 | 0 | 0 | 2 | 0 | 0 | 1 | 0 | 0 | 0 |
| March - April | 11 | 11 | 16 | 7 | 1 | 0 | 2 | 2 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |

We arbitrarily take 3 month groups and observe the mean, median and mode of these data (Table 17).
Table 17 Mean median and mode of the data in May-June, Sep - Oct and Nov-Dec

| Month | Mean | median | mode |
| :---: | :---: | :---: | :---: |
| May-June | 2.44 | 2 | 2 |
| Sep - Oct | 2.62 | 2 | 2 |
| Nov-Dec | 3.08 | 2 | 2 |

Since mean > mode and mean > median, the data exhibits positive skeweness with m.l.e. of $\theta>0$ and it is observed that the distribution is unimodal, hence HG distribution is suppose to give a better fit.

Table 18 Goodness of fit of HG distribution with $\alpha=0.05$

| Sample size $=50$ | MLE q $^{\wedge}$ | MLE $^{\wedge}$ | MLE k $^{\wedge}$ | Observed $\chi^{2}$ | d.f. | p-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| May-June | 0.4816 | 5.4108 | $2.1091(2)$ | 4.4928 | 4 | 0.3434 |
| Sep - Oct | 0.4936 | 5.0367 | $2.4050(2)$ | 1.9735 | 4 | 0.7406 |
| Nov-Dec | 0.5211 | 5.2204 | $2.3670(2)$ | 3.42 | 4 | 0.4901 |

Result: From the p-values it is seen that here HG distribution is a suitable.

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#### Abstract

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## BIOGRAPHY

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